# G22.3033-008, Spring 2010 Geometric Modeling

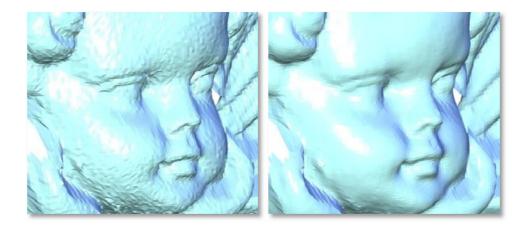
Linear algebra tools for geometric modeling

#### Least squares fitting

Motivation

- Why are we going over this again?
  - Many of the shape modeling methods presented in later lectures minimize functionals of the form

$$\mathbf{c}_{opt} = \underset{\mathbf{c}}{\operatorname{argmin}} \|\mathbf{A}\mathbf{c} - \mathbf{b}\|^2$$







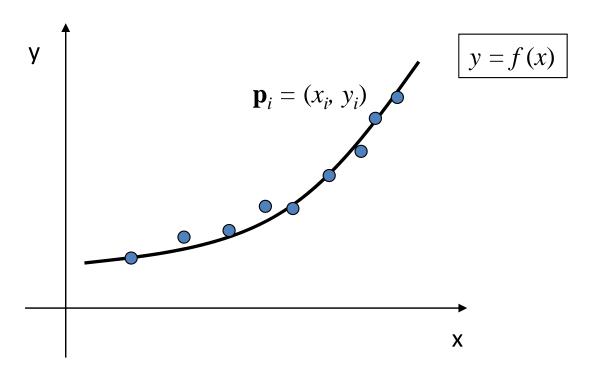




#### Least squares fitting

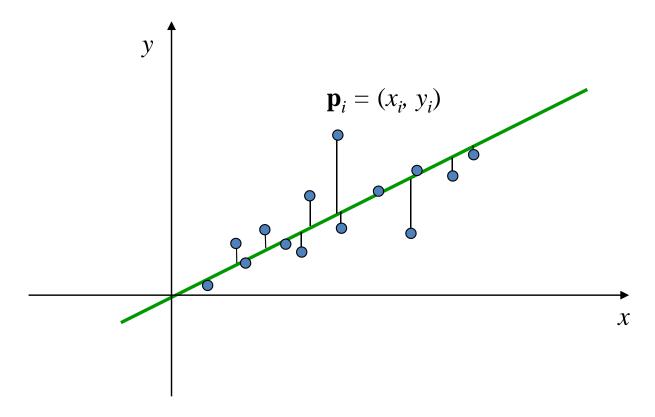
Motivation

 Given data points, fit a function that is "close" to the points



line fitting – 1st order polynomial in 2D

y-offsets minimization



line fitting – 1<sup>st</sup> order polynomial in 2D

• Find a line y = ax + b that minimizes

$$E(a,b) = \sum_{i=1}^{n} [y_i - (ax_i + b)]^2$$

- E(a,b) is quadratic in the unknown parameters a, b
- Another option would be, for example:

$$AbsErr(a,b) = \sum_{i=1}^{n} |y_i - (ax_i + b)|$$

■ But – it is not differentiable, harder to minimize...

line fitting – LS minimization

■ To find optimal a, b we differentiate E(a, b):

$$E(a,b) = \sum_{i=1}^{n} [y_i - (ax_i + b)]^2$$

$$\frac{\partial}{\partial a}E(a, b) = \sum_{i=1}^{n}(-2x_i)[y_i - (ax_i + b)] = 0$$

$$\frac{\partial}{\partial b}E(a, b) = \sum_{i=1}^{n} (-2)[y_i - (ax_i + b)] = 0$$

line fitting – LS minimization

• We obtain two linear equations for a, b:

$$\sum_{i=1}^{n} (-2x_i)[y_i - (ax_i + b)] = 0$$

$$\sum_{i=1}^{n} (-2)[y_i - (ax_i + b)] = 0$$

line fitting – LS minimization

• We get two linear equations for a, b:

(1) 
$$\sum_{i=1}^{n} \left[ x_i y_i - a x_i^2 - b x_i \right] = 0$$

(2) 
$$\sum_{i=1}^{n} [y_i - ax_i - b] = 0$$

line fitting – LS minimization

• We get two linear equations for a, b:

$$\left(\sum_{i=1}^{n} x_{i}^{2}\right) a + \left(\sum_{i=1}^{n} x_{i}\right) b = \sum_{i=1}^{n} x_{i} y_{i}$$

$$(\sum_{i=1}^{n} x_i) a + (\sum_{i=1}^{n} 1) b = \sum_{i=1}^{n} y_i$$

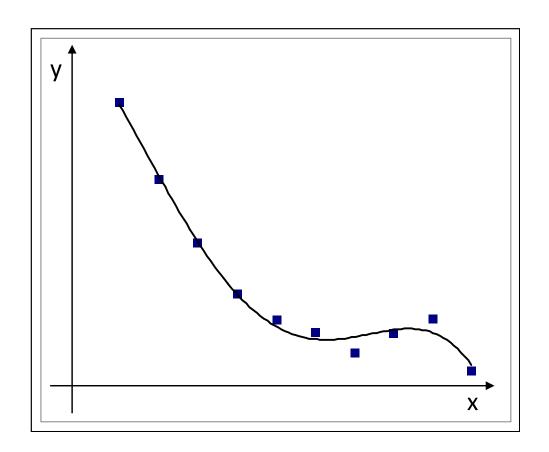
line fitting – LS minimization

Solve for a, b using e.g. Gauss elimination

• Question: why the solution is the *minimum* for the error function?

$$E(a, b) = \sum_{i=1}^{n} [y_i - (ax_i + b)]^2$$

## Fitting polynomials



## Fitting polynomials

- Decide on the degree of the polynomial, k
- Want to fit  $f(x) = a_k x^k + a_{k-1} x^{k-1} + ... + a_1 x + a_0$
- Minimize:

$$E(a_0, a_1, ..., a_k) = \sum_{i=1}^{n} [y_i - (a_k x_i^k + a_{k-1} x_i^{k-1} + ... + a_1 x_i + a_0)]^2$$

$$\frac{\partial}{\partial a_m} E(a_0, \dots, a_k) = \sum_{i=1}^n (-2x^m) [y_i - (a_k x_i^k + a_{k-1} x_i^{k-1} + \dots + a_0)] = 0$$

#### Fitting polynomials

• We get a linear system of k+1 equations in k+1 variables

$$\begin{pmatrix}
\sum_{i=1}^{n} 1 & \sum_{i=1}^{n} x_{i} & \cdots & \sum_{i=1}^{n} x_{i}^{k} \\
\sum_{i=1}^{n} x_{i} & \sum_{i=1}^{n} x_{i}^{2} & \cdots & \sum_{i=1}^{n} x_{i}^{k+1} \\
\vdots & \vdots & \ddots & \vdots \\
\sum_{i=1}^{n} x_{i}^{k} & \sum_{i=1}^{n} x_{i}^{k+1} & \cdots & \sum_{i=1}^{n} x_{i}^{2k}
\end{pmatrix}
\begin{pmatrix}
a_{0} \\
a_{1} \\
\vdots \\
\vdots \\
a_{k}
\end{pmatrix}
=
\begin{pmatrix}
\sum_{i=1}^{n} 1 \cdot y_{i} \\
a_{1} \\
\vdots \\
a_{k}
\end{pmatrix}$$

## General parametric fitting

- We can use this approach to fit any function  $f(\mathbf{x})$ 
  - Specified by parameters  $c_1$ ,  $c_2$ ,  $c_3$ , ...
  - The expression  $f(\mathbf{x})$  linearly depends on the parameters.

$$f(\mathbf{x}) = c_1 f_1(\mathbf{x}) + c_2 f_2(\mathbf{x}) + \dots + c_k f_k(\mathbf{x})$$

• Minimize – find best  $c_1$ ,  $c_2$ ,  $c_3$  ...:

$$\sum_{i=1}^{n} ||f(\mathbf{p}_i) - f_i||^2 = \sum_{i=1}^{n} ||\sum_{j=1}^{k} \mathbf{c}_j f_j(\mathbf{p}_i) - f_i||^2$$

- Let's look at the problem a little differently:
  - lacktriangle We have data points  $oldsymbol{\mathbf{p}}_i$  and desired function values  $f_i$
  - We would like :

$$\forall i = 1, ..., n$$
:  $f(\mathbf{p}_i) - f_i$ 

- Strict interpolation is in general not possible
  - In polynomials: n+1 points define a unique interpolation polynomial of degree n.
  - So, if we have 1000 points and want a cubic polynomial, we probably won't find it...

• We have an over-determined linear system  $n \times k$ :

$$f(\mathbf{p}_{1}) = c_{1} f_{1}(\mathbf{p}_{1}) + c_{2} f_{2}(\mathbf{p}_{1}) + \dots + c_{k} f_{k}(\mathbf{p}_{1}) = f_{1}$$

$$f(\mathbf{p}_{2}) = c_{1} f_{1}(\mathbf{p}_{2}) + c_{2} f_{2}(\mathbf{p}_{2}) + \dots + c_{k} f_{k}(\mathbf{p}_{2}) = f_{2}$$
...
$$f(\mathbf{p}_{n}) = c_{1} f_{1}(\mathbf{p}_{n}) + c_{2} f_{2}(\mathbf{p}_{n}) + \dots + c_{k} f_{k}(\mathbf{p}_{n}) = f_{n}$$

In matrix form:

$$\begin{pmatrix} f_1(\mathbf{p}_1) & f_2(\mathbf{p}_1) & \dots & f_k(\mathbf{p}_1) \\ f_1(\mathbf{p}_2) & f_2(\mathbf{p}_2) & \dots & f_k(\mathbf{p}_2) \\ & & & \dots & & \\ \vdots & & \vdots & & \vdots & & \\ f_1(\mathbf{p}_n) & f_2(\mathbf{p}_n) & \dots & f_k(\mathbf{p}_n) \end{pmatrix} \begin{pmatrix} c_1 \\ c_2 \\ \vdots \\ c_k \end{pmatrix} = \begin{pmatrix} f_1 \\ f_2 \\ \vdots \\ c_k \end{pmatrix}$$

In matrix form:

$$Ac = b$$

where  $A = (f_j(\mathbf{p}_i))_{i,j}$  is a rectangular  $n \times k$  matrix, n > k

$$\mathbf{c} = (c_1, c_2, ..., c_k)^{\mathrm{T}}$$
  $\mathbf{b} = (f_1, f_2, ..., f_n)^{\mathrm{T}}$ 

$$\mathbf{c}$$

- More constrains than variables no exact solutions generally exist
- We want to find something that is an "approximate solution":

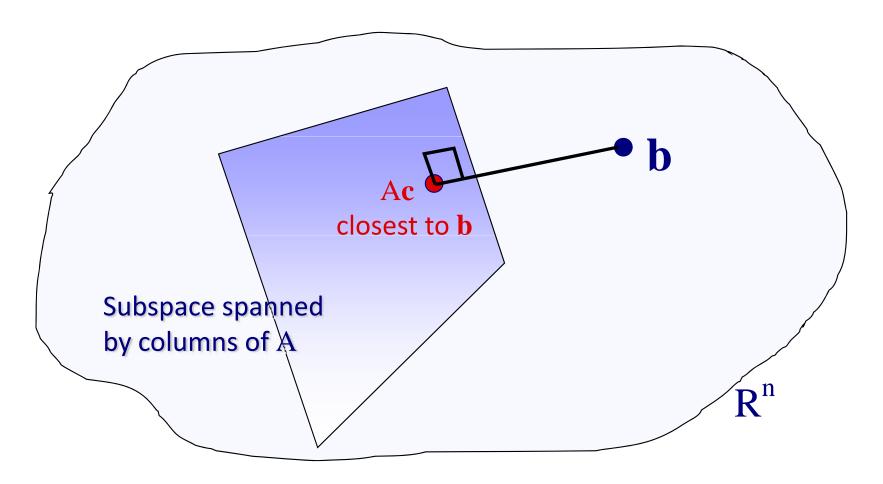
$$\mathbf{c}_{opt} = \underset{\mathbf{c}}{\operatorname{argmin}} \|\mathbf{A}\mathbf{c} - \mathbf{b}\|^2$$

- $\mathbf{c} \in \mathbb{R}^k$
- $Ac \in R^n$
- As we vary c, Ac varies over the linear subspace of R<sup>n</sup> spanned by the columns of A:

$$\mathbf{Ac} = \left( \begin{array}{c|c} A_1 & A_2 & A_k \end{array} \right) \left( \begin{array}{c} c_1 \\ c_2 \\ \vdots \\ c_k \end{array} \right) = c_1 \left( \begin{array}{c|c} A_1 & + c_2 \\ A_2 & + \dots + c_k \end{array} \right) + c_1 \left( \begin{array}{c|c} A_1 & + c_2 \\ \vdots & \vdots & \vdots \\ c_k & A_k \end{array} \right)$$

This is also known as the column space of A

• We want to find the closest  $\mathbf{Ac}$  to  $\mathbf{b}$ :  $\min_{\mathbf{c}} \|\mathbf{Ac} - \mathbf{b}\|^2$ 



■ The point Ac closest to b satisfies:

 $(Ac - b) \perp \{subspace of A's columns\}$ 



$$\forall \operatorname{column} A_i$$
:  $\langle A_i, \operatorname{A}\mathbf{c} - \mathbf{b} \rangle = 0$ 

$$\forall i, A_i^{\mathrm{T}}(\mathbf{Ac} - \mathbf{b}) = 0$$

These are called the normal equations

$$\frac{\mathbf{A}^{\mathrm{T}}(\mathbf{A}\mathbf{c} - \mathbf{b}) = 0}{(\mathbf{A}^{\mathrm{T}}\mathbf{A})\mathbf{c} = \mathbf{A}^{\mathrm{T}}\mathbf{b}}$$

• We have a square symmetric system  $(\mathbf{A}^{T}\mathbf{A})\mathbf{c} = \mathbf{A}^{T}\mathbf{b}$ 

$$(k \times k)$$

If A has full rank (the columns of A are linearly independent) then (A<sup>T</sup>A) is invertible.

$$\min_{\mathbf{c}} \|\mathbf{A}\mathbf{c} - \mathbf{b}\|^{2}$$

$$\downarrow \downarrow$$

$$\mathbf{c} = (\mathbf{A}^{T} \mathbf{A})^{-1} \mathbf{A}^{T} \mathbf{b}$$

#### Weighted least squares

If each constraint has a weight in the energy:

$$\min_{\mathbf{c}} \sum_{i=1}^{n} w_i \left( f_{\mathbf{c}}(\mathbf{p}_i) - f_i \right)^2$$

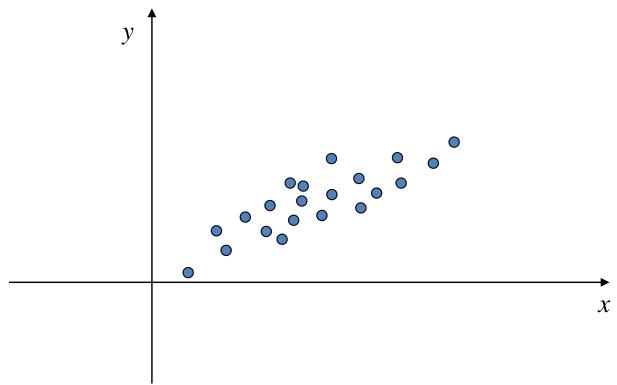
- The weights  $w_i > 0$  and don't depend on  $\mathbf{c}$
- Then:

min 
$$(\mathbf{Ac} - \mathbf{b})^{\mathrm{T}} \mathbf{W}^{\mathrm{T}} \mathbf{W} (\mathbf{Ac} - \mathbf{b})$$
 where  $\mathbf{W} = (w_i)_{ii}$ 

$$(A^T W^2 A) \mathbf{c} = A^T W^2 \mathbf{b}$$

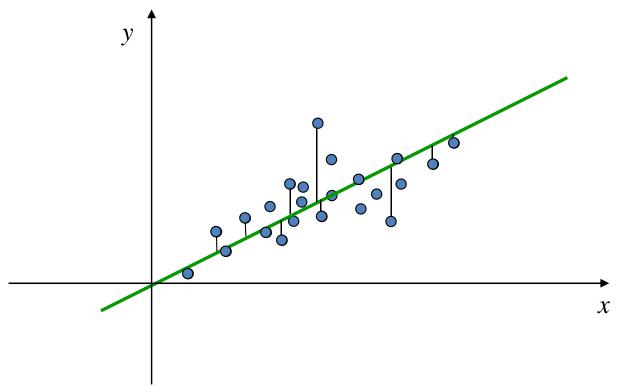
But first, reminder about eigenvectors and eigenvalues

#### Motivation



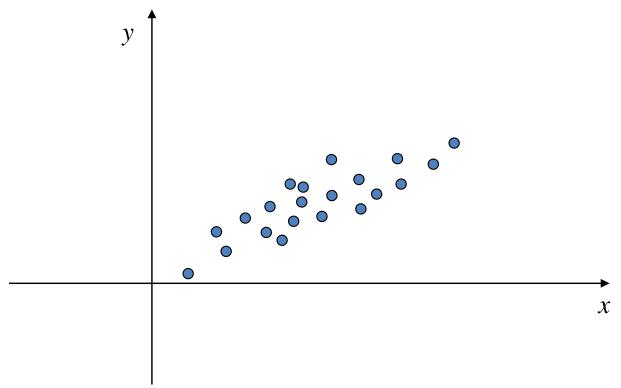
 Given a set of points, find the best line that approximates them

#### Motivation

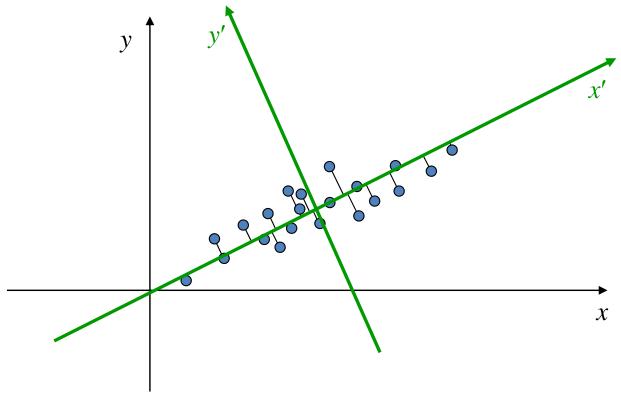


• We just saw how to fit a parametric line y = ax + b, but this does not work for vertical lines

#### Motivation



How to fit a line such that the true (orthogonal) distances are minimized?

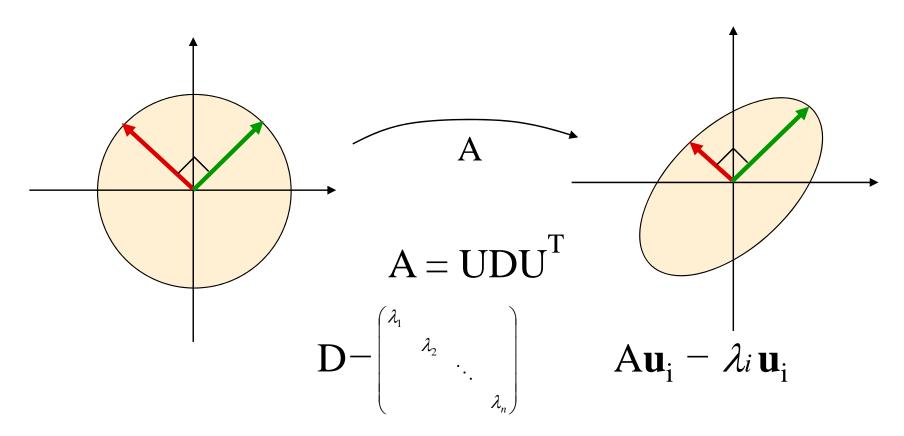


PCA finds axes that minimize the sum of distances<sup>2</sup>

#### Linear algebra recap

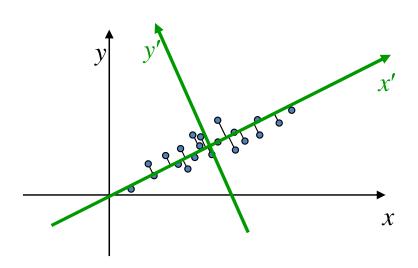
Symmetric matrices

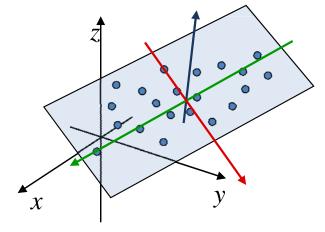
 If A is symmetric, the eigenvectors are orthogonal and there's always an eigenbasis.



Basic idea

 PCA finds an orthogonal basis that best represents given data set

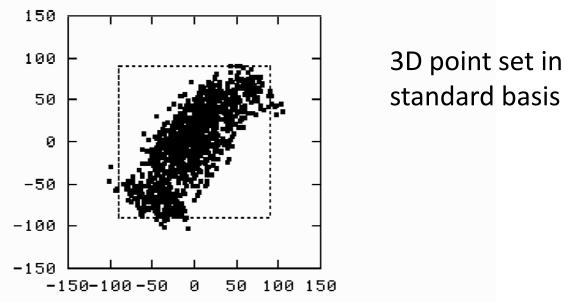




■ PCA finds a best approximating line/plane/axes... (in terms of  $\Sigma_{distances}^2$ )

Basic idea

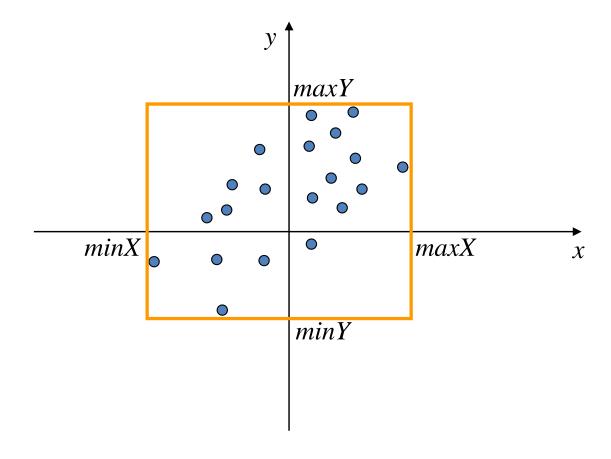
 PCA finds an orthogonal basis that best represents given data set



■ PCA finds a best approximating line/plane/axes... (in terms of  $\Sigma_{distances}^2$ )

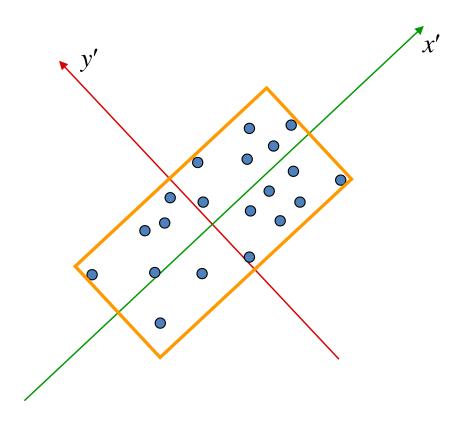
**Applications** 

An axis-aligned bounding box: agrees with the standard axes



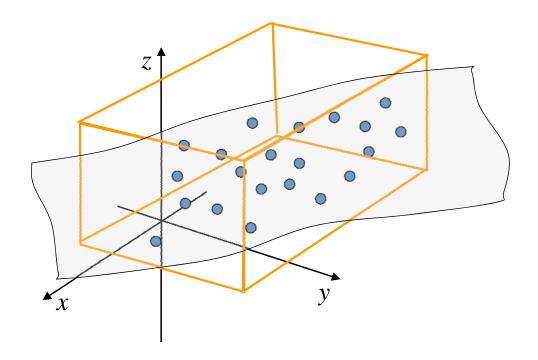
Application: oriented bounding box

Tighter fit



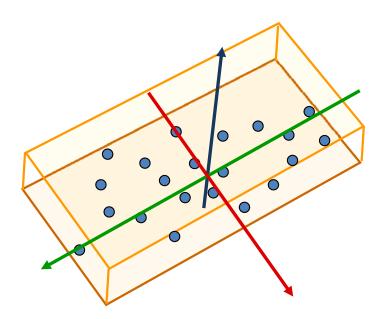
Application: oriented bounding box

Axis aligned bounding box



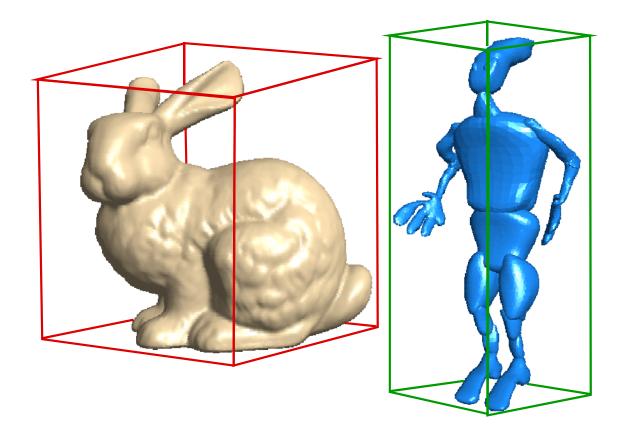
Application: oriented bounding box

Oriented bounding box by PCA

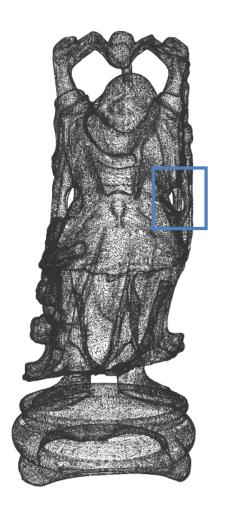


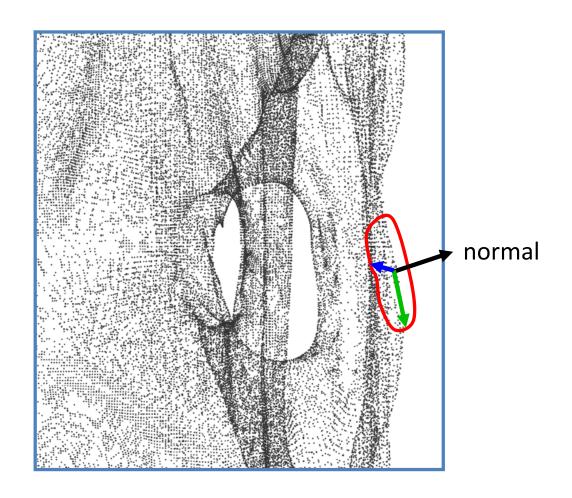
Application: oriented bounding box

- Serve as very simple "approximation" of the object
- Fast collision detection, visibility queries
- Whenever we need to know the dimensions (size) of the object
- The models consist of thousands of polygons
- To quickly test that they don't intersect, the bounding boxes are tested
- Sometimes a hierarchy of BB's is used
- The tighter the BB the less "false alarms" we have

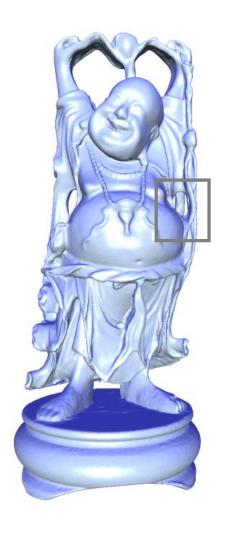


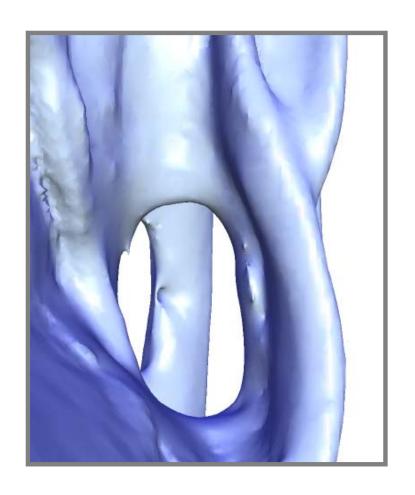
Application: local frame fitting





Application: estimate normals





Application: shape alignment

3D search engines (see http://shape.cs.princeton.edu/)

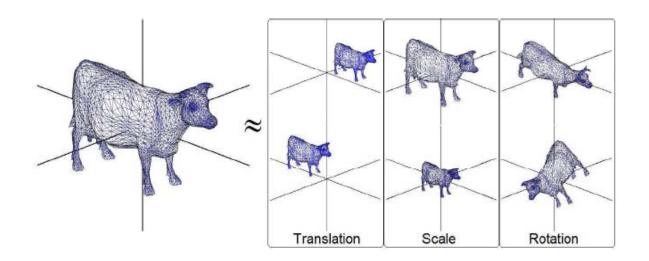


Query



Application: shape alignment

 Can use PCA to find canonical axes and scale for shape comparison

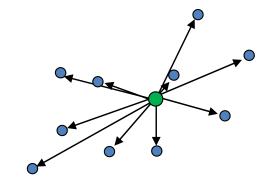


#### **Notations**

■ Denote our data points by  $\mathbf{x}_1$ ,  $\mathbf{x}_2$ , ...,  $\mathbf{x}_n \in R^d$ 

Center of mass:

$$\mathbf{m} = \frac{1}{n} \sum_{i=1}^{n} \mathbf{x}_{i}$$



Vectors from the centroid:

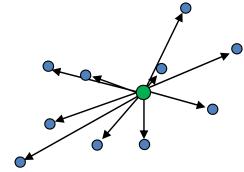
$$\mathbf{y}_i = \mathbf{x}_i - \mathbf{m}$$

# The origin of the new axes

 The origin of the new axes will be the center of mass m

It can be shown that:

$$\mathbf{m} = \underset{\mathbf{x}}{\operatorname{argmin}} \sum_{i=1}^{n} \|\mathbf{x}_{i} - \mathbf{x}\|^{2}$$



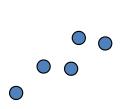
$$\mathbf{m} = \frac{1}{n} \sum_{i=1}^{n} \mathbf{x}_{i}$$

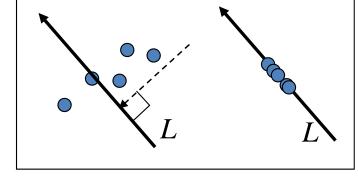
# Variance of projected points

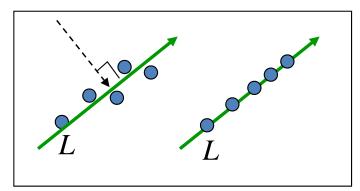
- Let us measure the variance (scatter) of our points in different directions
- Let's look at a line L through the center of mass  $\mathbf{m}$ , and project our points  $\mathbf{x}_i$  onto it. The variance of the projected points  $\mathbf{x}'_i$  is:

Want to find directions of maximal/minimal variance

$$\operatorname{var}(L) = \frac{1}{n} \sum_{i=1}^{n} ||\mathbf{x}'_i - \mathbf{m}||^2$$







Original set

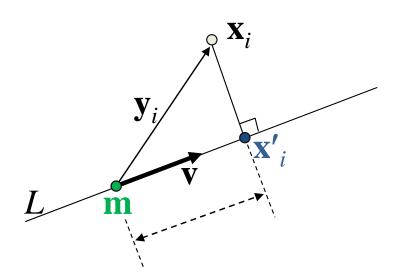
Small variance

Large variance

# Variance of projected points

- Given a direction  $\mathbf{v}$ ,  $||\mathbf{v}|| = 1$
- Line L through  $\mathbf{m}$  in the direction of  $\mathbf{v}$  is  $L(t) = \mathbf{m} + \mathbf{v}t$ .

$$||\mathbf{x'}_i - \mathbf{m}|| = \langle \mathbf{v}, \mathbf{x}_i - \mathbf{m} \rangle / ||\mathbf{v}|| = \langle \mathbf{v}, \mathbf{y}_i \rangle = \mathbf{v}^{\mathrm{T}} \mathbf{y}_i = \mathbf{y}_i^{\mathrm{T}} \mathbf{v}$$



# Variance of projected points

So,  

$$\operatorname{var}(L) = \frac{1}{n} \sum_{i=1}^{n} \|\mathbf{x}_{i}' - \mathbf{m}\|^{2} = \frac{1}{n} \sum_{i=1}^{n} (\mathbf{y}_{i}^{\mathrm{T}} \mathbf{v})^{2} = \frac{1}{n} \|\mathbf{Y}^{\mathrm{T}} \mathbf{v}\|^{2} = \frac{1}{n} (\mathbf{Y}^{\mathrm{T}} \mathbf{v})^{\mathrm{T}} (\mathbf{Y}^{\mathrm{T}} \mathbf{v}) = \frac{1}{n} \mathbf{v}^{\mathrm{T}} \mathbf{Y} \mathbf{Y}^{\mathrm{T}} \mathbf{v} = \mathbf{v}^{\mathrm{T}} S \mathbf{v}.$$

$$S = (1/n) Y Y^T$$
 Scatter matrix

where Y is a  $d \times n$  matrix with  $\mathbf{y}_k = \mathbf{x}_k - \mathbf{m}$  as columns.

 $\blacksquare$  The scatter matrix S measures the variance of our points

### Directions of maximal variance

- So, we have:  $var(L) = \mathbf{v}^{T} \mathbf{S} \mathbf{v}$
- Theorem:

Let 
$$f: \{\mathbf{v} \in R^d \mid ||\mathbf{v}|| = 1\} \to R$$
,  $f(\mathbf{v}) = \mathbf{v}^T S \mathbf{v}$  (and  $S$  is a symmetric matrix).

Then, the extrema of f are attained at the eigenvectors of S.

So, eigenvectors of S are directions of maximal/minimal variance!

#### Directions of maximal variance

- Find extrema of  $\mathbf{v}^{\mathrm{T}} \mathbf{S} \mathbf{v}$
- side condition  $\mathbf{v}^{\mathrm{T}}\mathbf{v}=1$
- Lagrange Multipliers:  $\nabla f + \lambda \nabla g = 0$

$$\nabla (\mathbf{v}^{\mathrm{T}} \mathbf{S} \mathbf{v}) + \lambda \nabla (\mathbf{v}^{\mathrm{T}} \mathbf{v} - 1) = 0$$

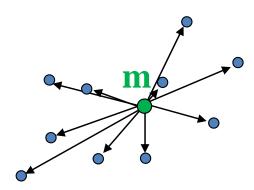
$$\mathbf{S} \mathbf{v} + \lambda \mathbf{v} = 0$$

$$\mathbf{S} \mathbf{v} = -\lambda \mathbf{v}$$

This is the definition of an eigenvector of S

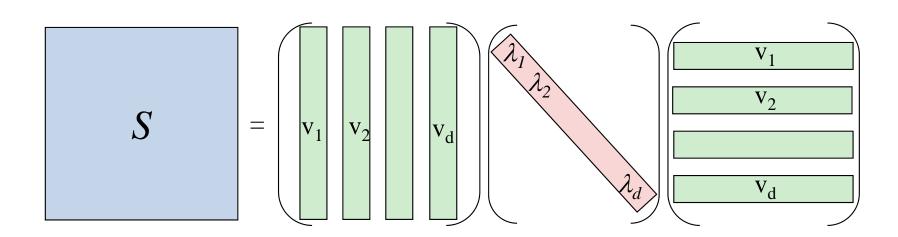
# Summary so far

- We take the centered data vectors  $\mathbf{y}_1$ ,  $\mathbf{y}_2$ , ...,  $\mathbf{y}_n \in R^d$
- Construct the scatter matrix  $S = Y Y^T$
- S measures the variance of the data points
- Eigenvectors of S are directions of maximal variance.



# Scatter matrix eigendecomposition

- S is symmetric
- $\Rightarrow$  S has eigendecomposition:  $S = VDV^T$



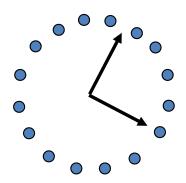
The eigenvectors form orthogonal basis

## Principal components

- Eigenvectors that correspond to big eigenvalues are the directions in which the data has strong components (= large variance).
- If the eigenvalues are more or less the same –
   there is no preferable direction.

Note: the eigenvalues are always nonnegative. Think why...

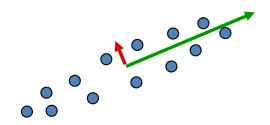
## Principal components



- There's no preferable direction
- S looks like this:

$$\mathbf{V} \begin{pmatrix} \lambda & \\ & \lambda \end{pmatrix} \mathbf{V}^{\mathrm{T}}$$

Any vector is an eigenvector



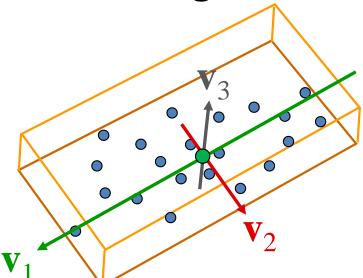
- There's a clear preferable direction
- S looks like this:

$$\mathbf{V} \begin{pmatrix} \lambda & \\ & \mu \end{pmatrix} \mathbf{V}^{\mathrm{T}}$$

•  $\mu$  is close to zero, much smaller than  $\lambda$ 

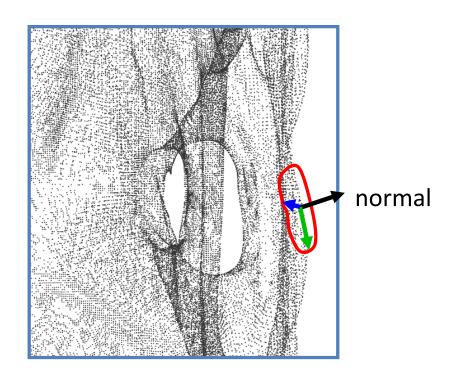
Oriented bounding box

■ For finding oriented bounding box or alignment — we simply compute the bounding box with respect to the axes defined by the eigenvectors. The origin is at the centroid **m**.

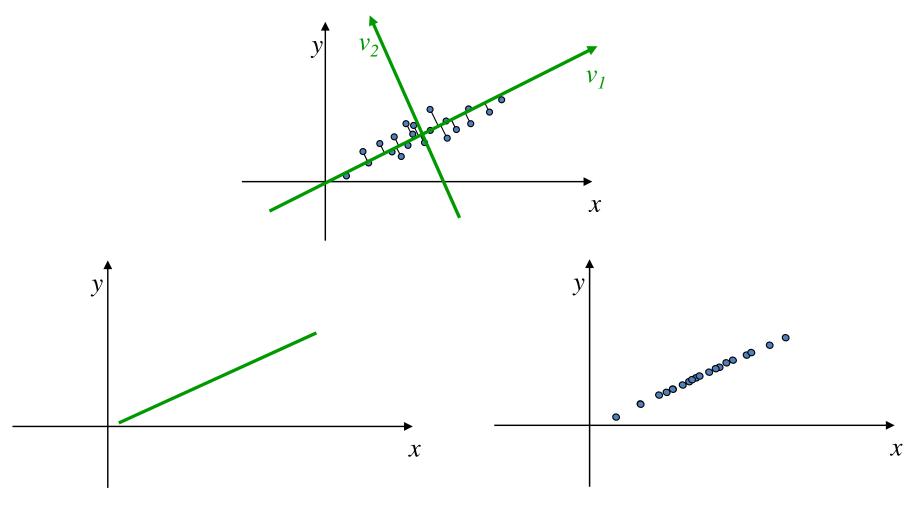


Local frame/normal estimation

- Sort the eigenvectors by ascending eigenvalues
- The eigenvector with  $\lambda \approx 0$  is the normal



Dimensionality reduction / approximation

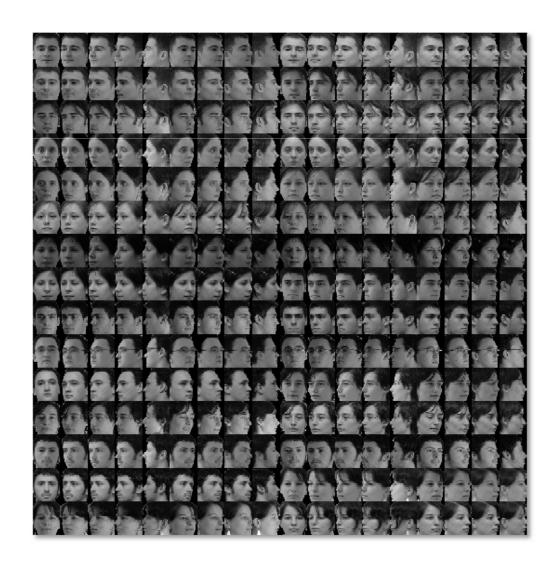


This line segment approximates the original data set

The projected data set approximates the original data set

Dimensionality reduction / approximation

- Each image is 64x64
- Vector in R<sup>64x64</sup>
- But in fact all the faces live on a lowdimensional subspace
- Can find meaningful axes with PCA and other methods
  - face pose
  - expression
  - **-** ...



Dimensionality reduction / approximation

• In general dimension d, the eigenvalues are sorted in descending order:

$$\lambda_1 \ge \lambda_2 \ge \dots \ge \lambda_d$$

- The eigenvectors are sorted accordingly.
- To get an approximation of dimension d' < d, we take the d' first eigenvectors and look at the subspace they span (d' = 1 is a line, d' = 2 is a plane...)

Dimensionality reduction / approximation

To get an approximating set, we project the original data points onto the chosen subspace:

$$\mathbf{x}_i = \mathbf{m} + \alpha_1 \mathbf{v}_1 + \alpha_2 \mathbf{v}_2 + \dots + \alpha_{d'} \mathbf{v}_{d'} + \dots + \alpha_d \mathbf{v}_d$$

Projection:

$$\mathbf{x}_{i}' = \mathbf{m} + \alpha_{1}\mathbf{v}_{1} + \alpha_{2}\mathbf{v}_{2} + \dots + \alpha_{d'}\mathbf{v}_{d'} + \mathbf{0}\cdot\mathbf{v}_{d'+1} + \dots + \mathbf{0}\cdot\mathbf{v}_{d}$$

#### Technical remarks:

- $\lambda_i \geq 0$ , i = 1,...,d (such matrices are called positive semi-definite). So we can indeed sort by the magnitude of  $\lambda_i$
- Theorem:  $\lambda_i \geq 0 \iff \langle \mathbf{S}\mathbf{v}, \mathbf{v} \rangle \geq 0 \quad \forall \mathbf{v}$ Proof:  $\mathbf{S} = \mathbf{V}\mathbf{D}\mathbf{V}^T \Rightarrow \langle \mathbf{S}\mathbf{v}, \mathbf{v} \rangle = \mathbf{v}^T\mathbf{S}\mathbf{v} = \mathbf{v}^T\mathbf{V}\mathbf{D}\mathbf{V}^T\mathbf{v} = (\mathbf{V}^T\mathbf{v})^T\mathbf{D}(\mathbf{V}^T\mathbf{v}) = \mathbf{w}^T\mathbf{D}\mathbf{w} = 2 \mathbf{v}^T\mathbf{v} + 2 \mathbf{v}^T\mathbf{v} +$

Therefore,  $\lambda_i \geq 0 \iff \langle \mathbf{S}\mathbf{v}, \mathbf{v} \rangle \geq 0 \quad \forall \mathbf{v}$